

# Generalized pairwise comparisons for precision medicine

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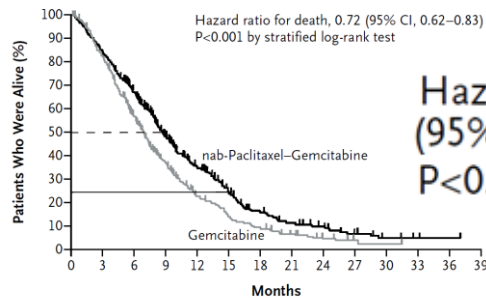
- *Evidence-based medicine*
  - (Meta-analyses of) randomized control trials
  - Subgroup analyses, if appropriate
- *Precision medicine*

“Giving the right treatment to the right patient at the right time”
- *Personalized medicine*
  - Precision medicine with personalized/patient-centric choices for therapeutic decisions

# An Unmet Statistical Need

Consider the following results

rival



No. at Risk	0	3	6	9	12	15	18	21	24	27	30	33	36	39
nab-Paclitaxel-Gemcitabine	431	357	269	169	108	67	40	27	16	9	4	1	1	0
Gemcitabine	430	340	220	124	69	40	26	15	7	3	1	0	0	0

	Worst grade related AE	Monotherapy (n=430)	Combination (n=431)
Grade 3			
Grade 4		<b>23%</b>	<b>54%</b>

A patient might reason:

- Taking combination, I’m more likely to live longer (by how much?)
- Taking combination, I’m more likely to have grade 3/4 adverse events (AEs)
- I’m willing to experience AEs for a survival benefit of at least  $m$  months...

- A single (primary) endpoint drives decision-making
- Other endpoints are analyzed descriptively
- Safety informally balanced against efficacy, resulting in debatable risk / benefit analyses
- Patient preferences are not formally taken into account

Research Article

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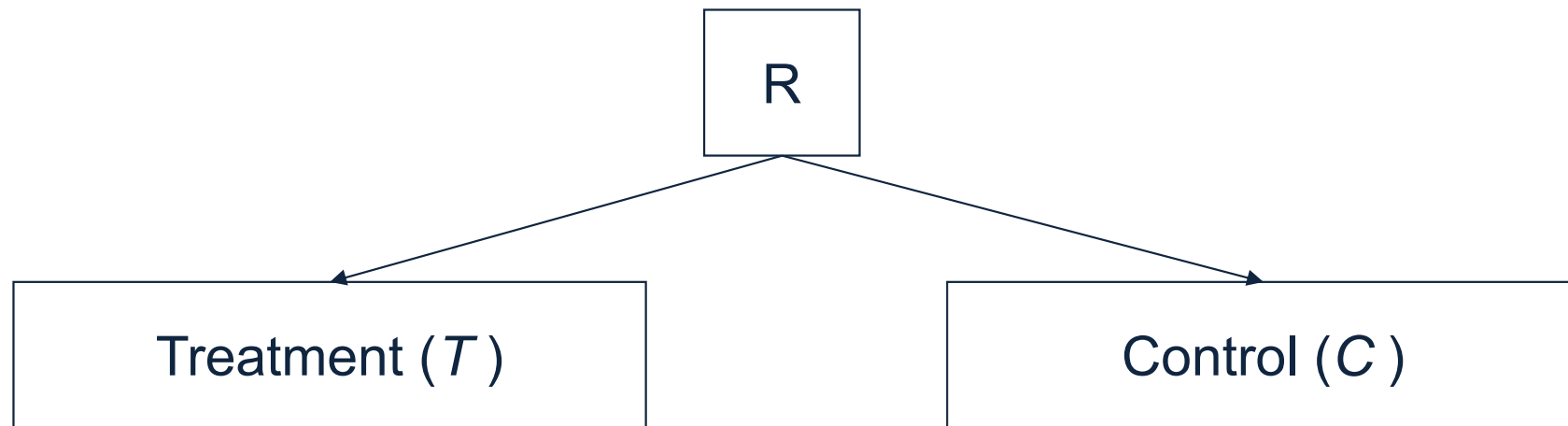
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# Generalized pairwise comparisons of prioritized outcomes in the two-sample problem

Marc Buyse<sup>a,b\*†</sup>

# Randomized Trial

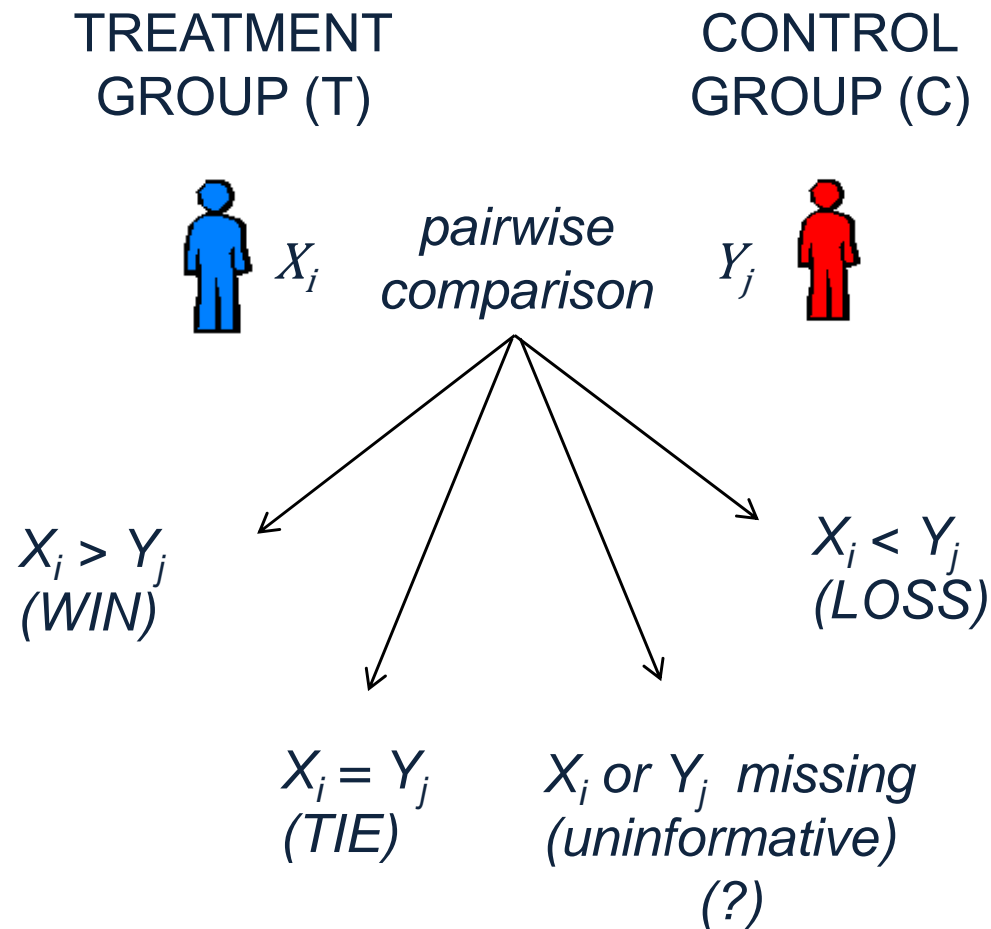


Let  $X_i$  be the outcome of  
 $i^{th}$  subject in  $T$  ( $i = 1, \dots, n$ )

Let  $Y_j$  be the outcome of  
 $j^{th}$  subject in  $C$  ( $j = 1, \dots, m$ )

# Pairwise Comparisons

Let  $X_i$  and  $Y_j$  be the observed values of a continuous outcome



- Let  $p_{ij}$  be equal to
  - 1 if the pair is a win
  - 1 if the pair is a loss
  - 0 if the pair is a tie/?

- Then

$$U = \sum_{ij} p_{ij} / (nm)$$

- A generalization of the Wilcoxon-Mann-Whitney test-statistic

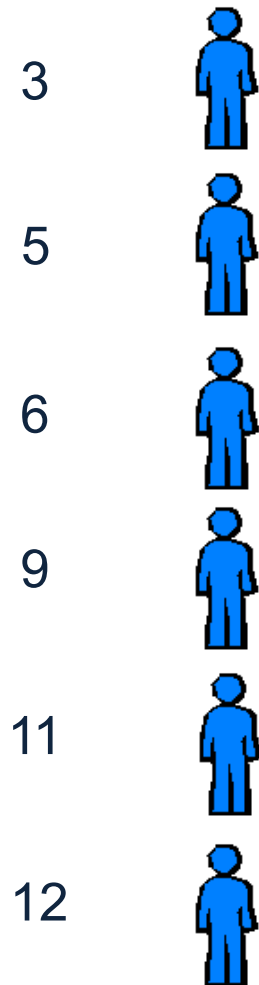
- And (if no missing data)

$$\Delta = E(U) = P(X > Y) - P(X < Y) = P(T \text{ better}) - P(C \text{ better})$$

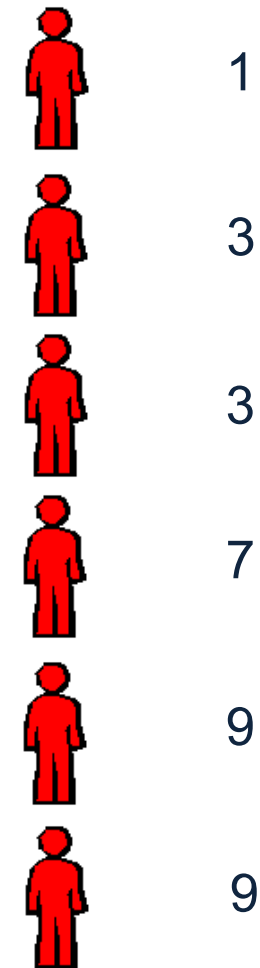


# Illustration

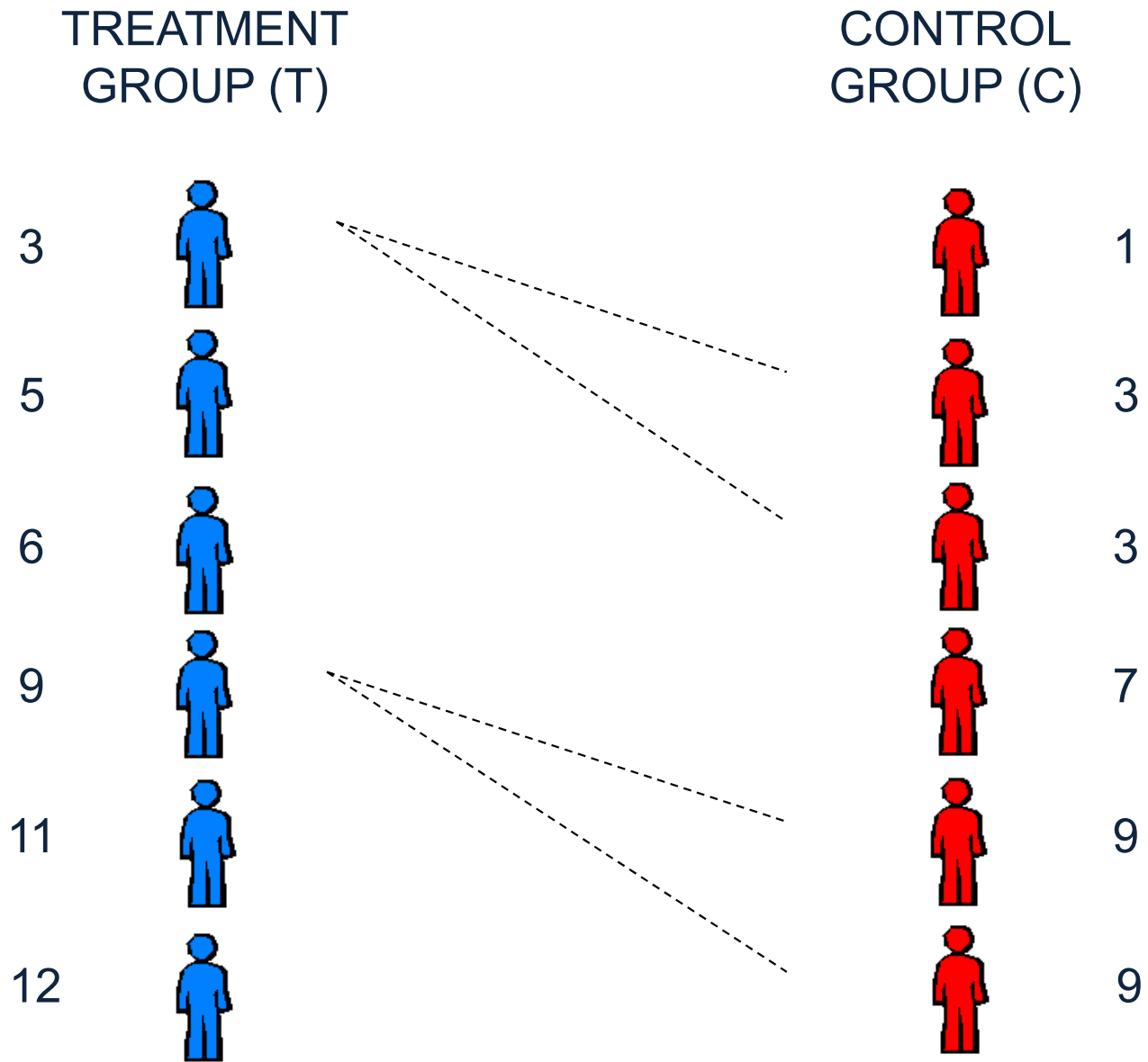
TREATMENT  
GROUP (T)



CONTROL  
GROUP (C)



# Ties

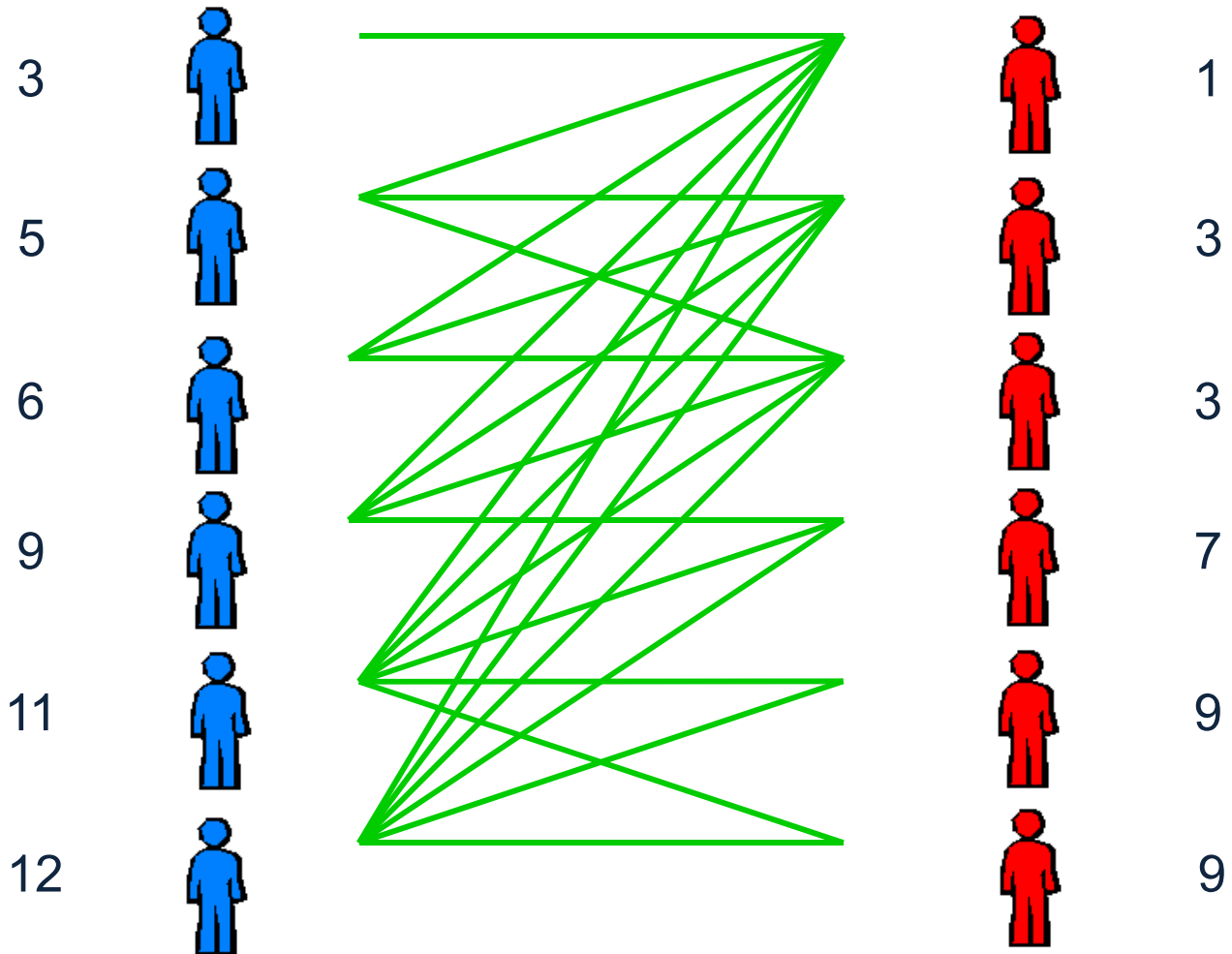


4 TIES

# Wins

TREATMENT  
GROUP (T)

CONTROL  
GROUP (C)

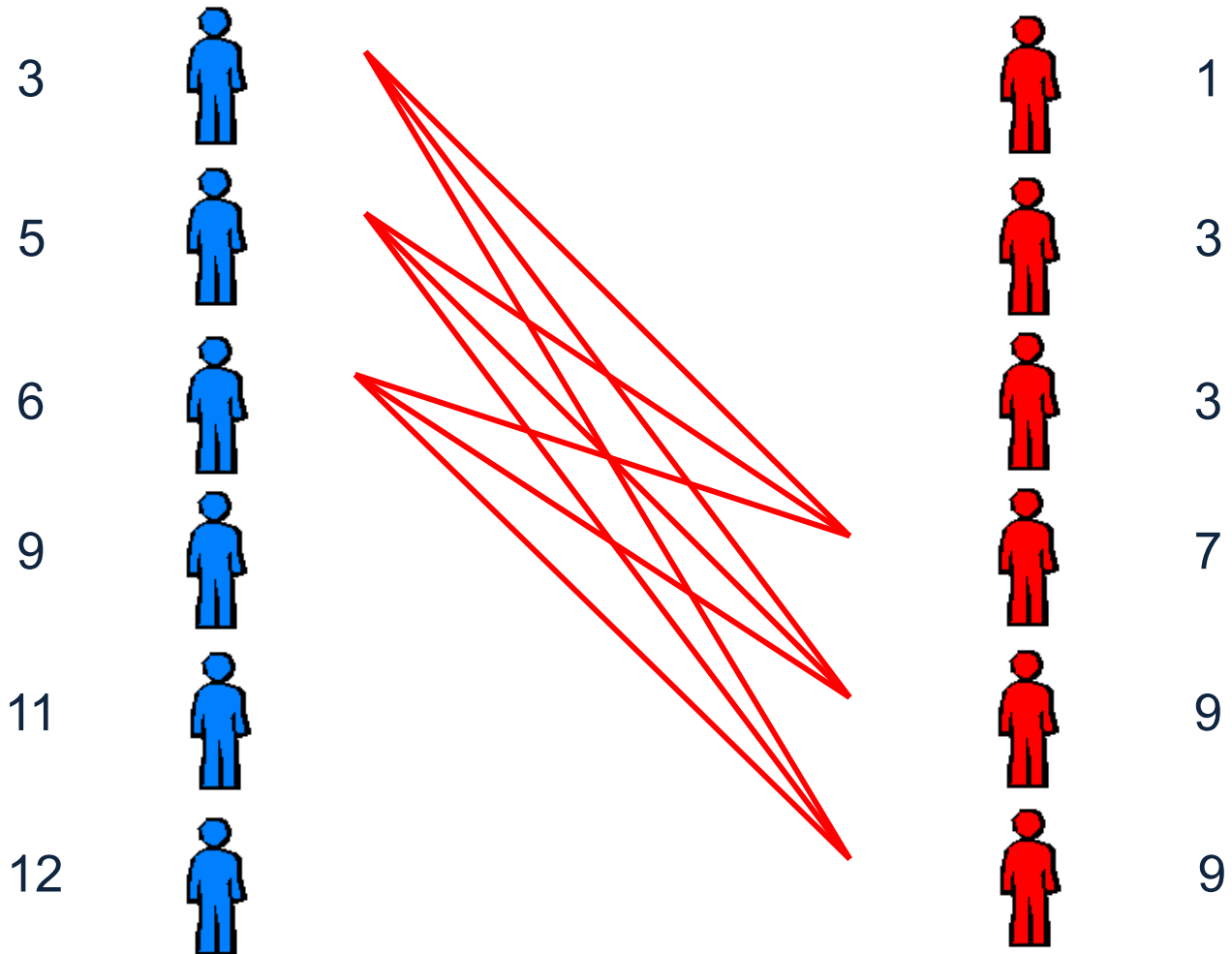


23 WINS

# Losses

TREATMENT  
GROUP (T)

CONTROL  
GROUP (C)



9 LOSSES

# Net Benefit

Ties	Wins	Losses	Net benefit
$4 / 36 = 0.11$	$23 / 36 = 0.64$	$9 / 36 = 0.25$	$0.64 - 0.25 = 0.39$

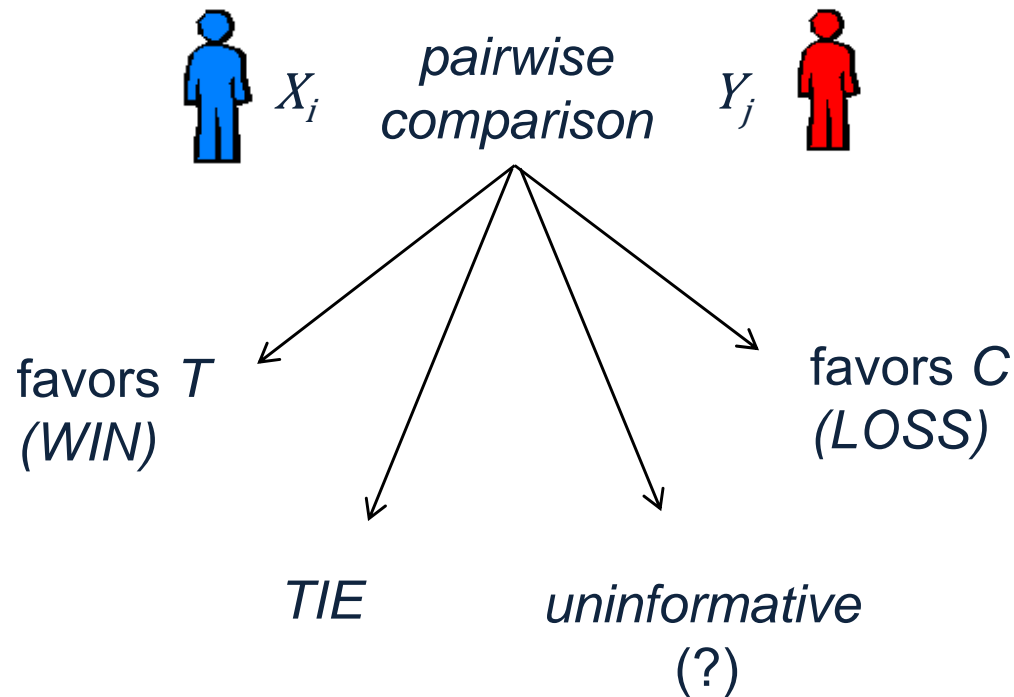
The probability of a patient having a better outcome

- if on treatment is 0.64
- if on control is 0.25

The "net benefit" is 0.39

*Note: the "win ratio" is  $0.64 / 0.25 = 2.56$*

Now let  $X_i$  and  $Y_j$  be the observed values of any outcome measure (continuous, time-to-event, binary, categorical, ...)



# Time to Event

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$X_i$ censored	$Y_j$ censored	$X_i > Y_j$	$X_i < Y_j$	$X_i = Y_j$
No	No	Win	Loss	Tie
Yes	No	Win	?	?
No	Yes	?	Loss	?
Yes	Yes	?	?	?

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# Thresholds of Clinical Relevance

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$X_i$ censored	$Y_j$ censored	$X_i - Y_j > m$	$X_i - Y_j < -m$	$ X_i - Y_j  \leq m$
No	No	Win	Loss	Tie
Yes	No	Win	?	?
No	Yes	?	Loss	?
Yes	Yes	?	?	?

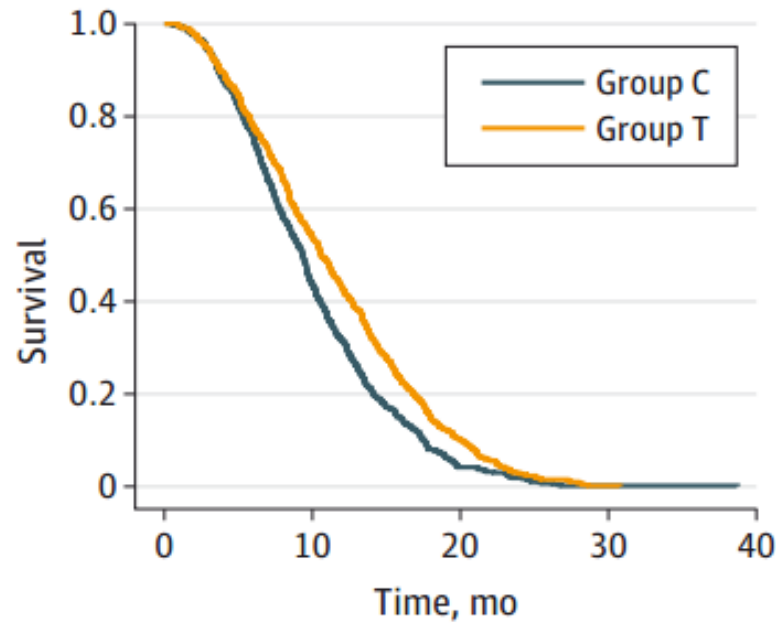
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# Net Benefit – Proportional Hazards

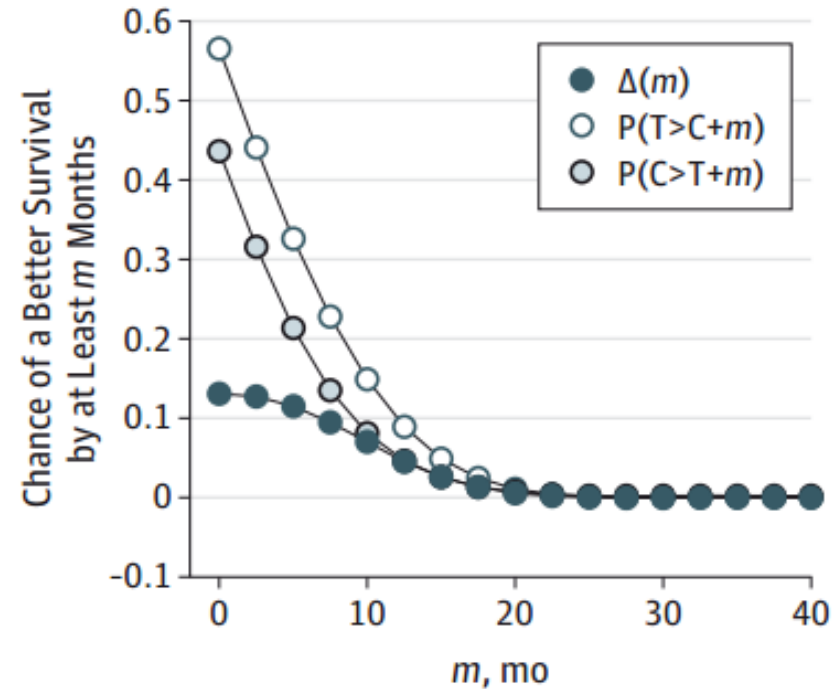
KAPLAN-MEIER CURVES

**A** Scenario 1: proportional hazards



No. at risk		0	10	20	30
Group C	600	263	26	1	
Group T	600	324	62	1	

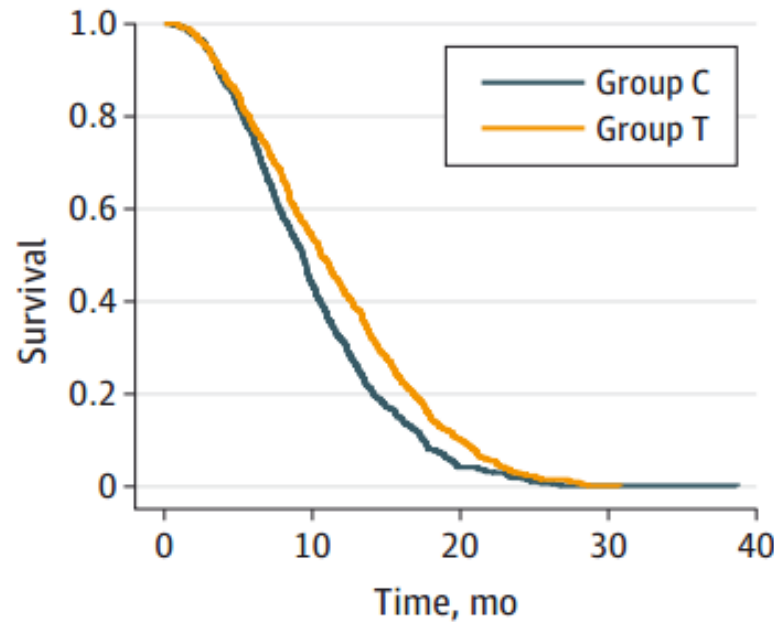
NET BENEFIT OF AT LEAST  $m$  MONTHS



# Net Benefit – Proportional Hazards

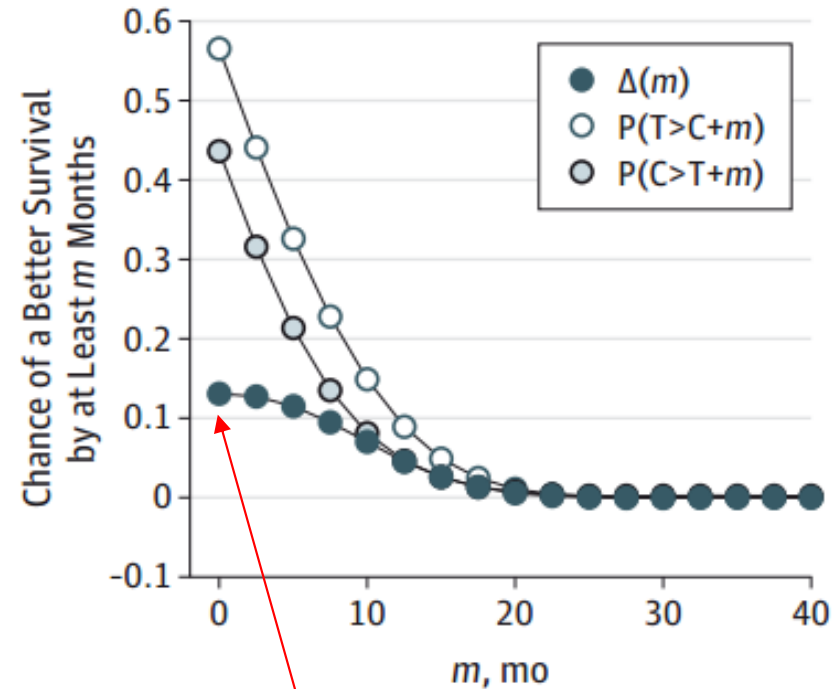
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NET BENEFIT OF AT LEAST  $m$  MONTHS



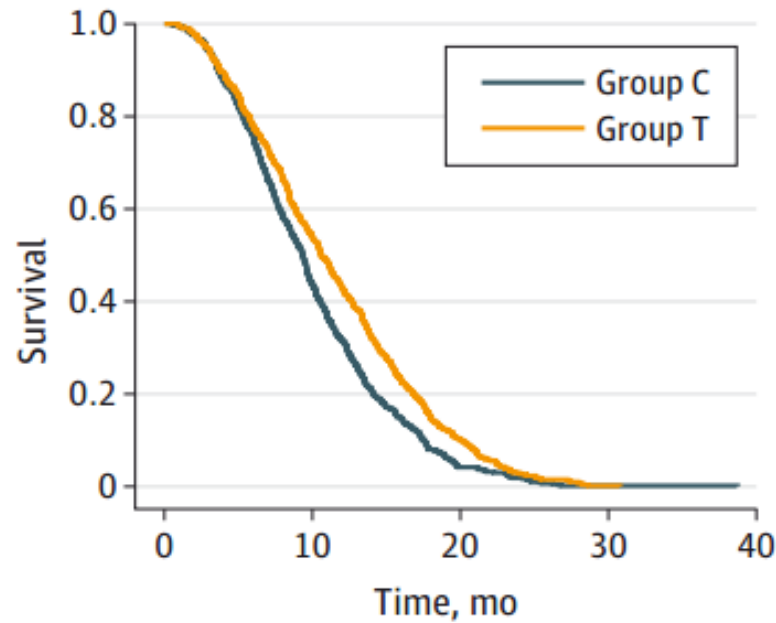
*There is a 13% net probability that survival will be longer on T than C*

*The “net benefit” of T is 13%*

# Net Benefit – Proportional Hazards

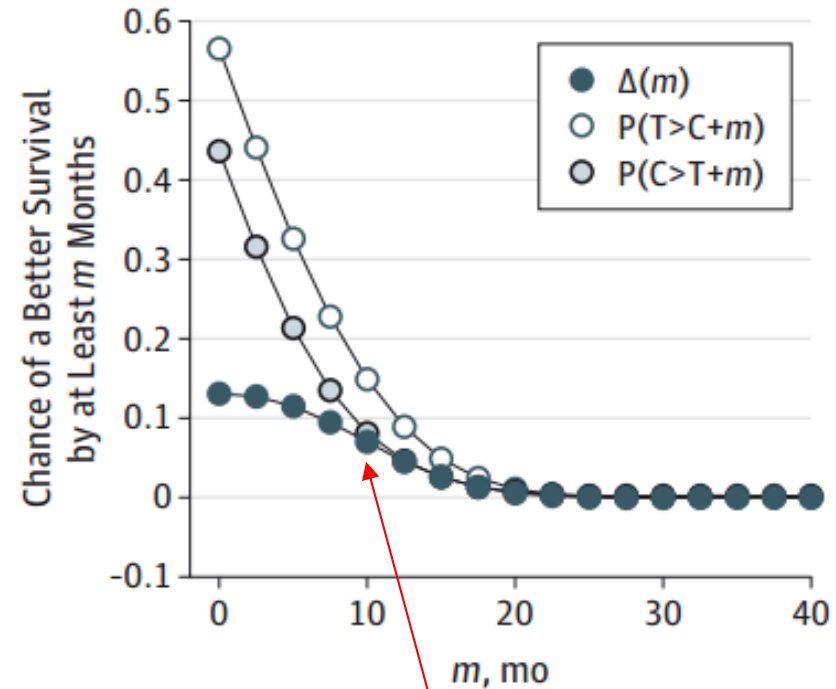
KAPLAN-MEIER CURVES

**A** Scenario 1: proportional hazards



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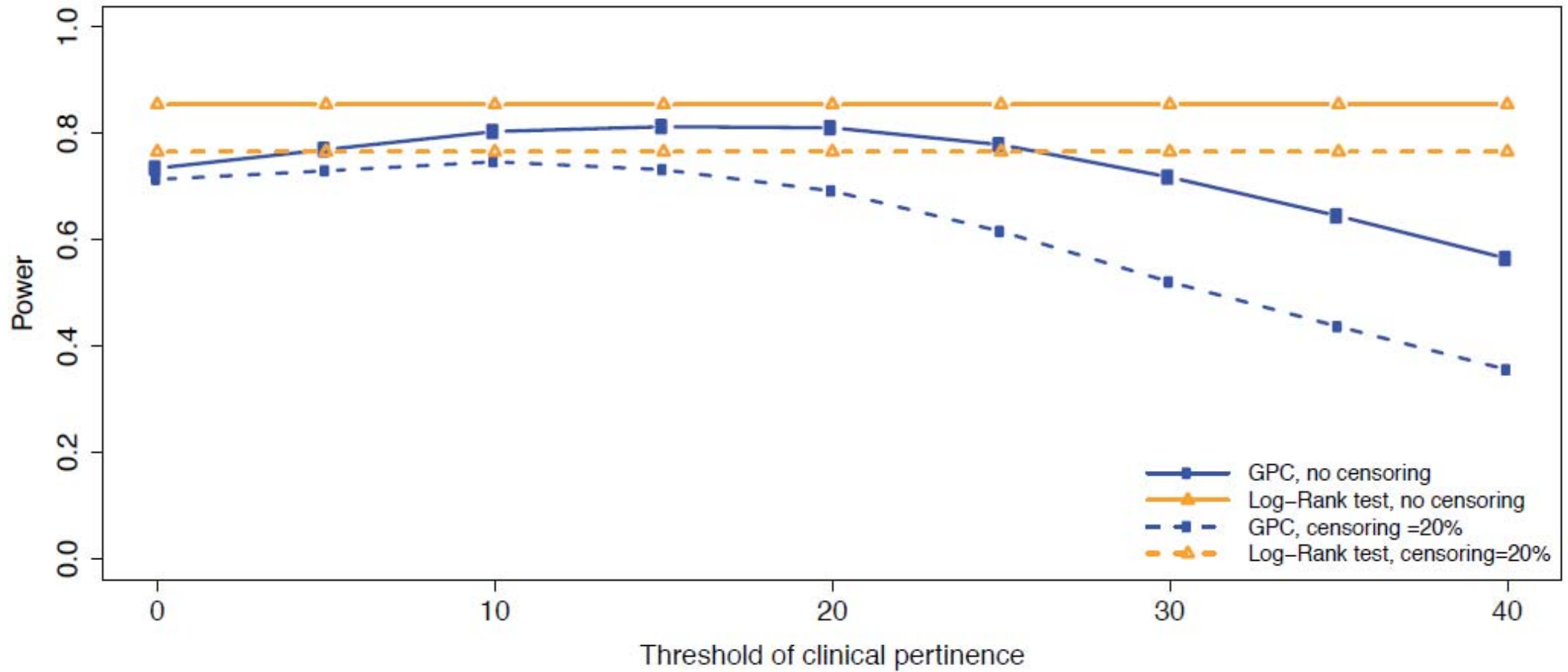
NET BENEFIT OF AT LEAST  $m$  MONTHS



*There is an 8% net benefit of at least 10 months in favor of T*

# Power – Proportional Hazards

Power of several tests in the proportional hazards scenario

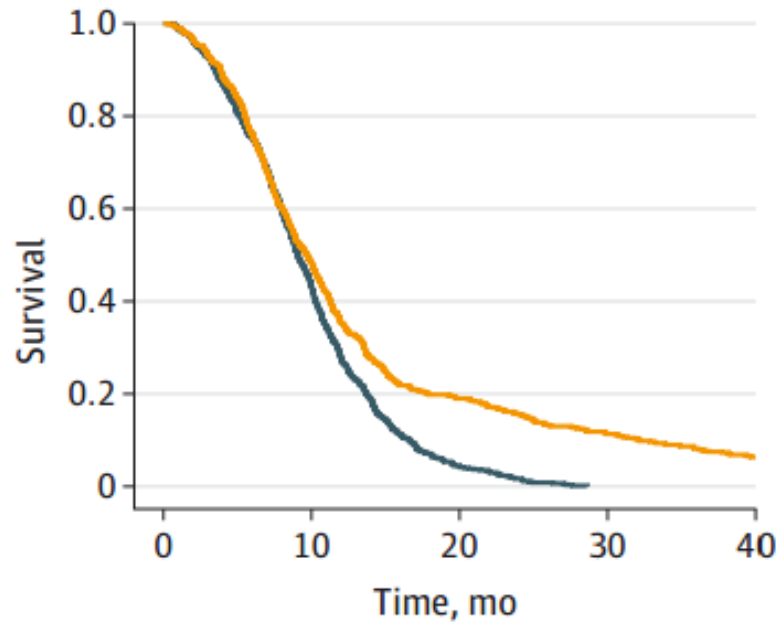


# Net Benefit – Delayed Difference

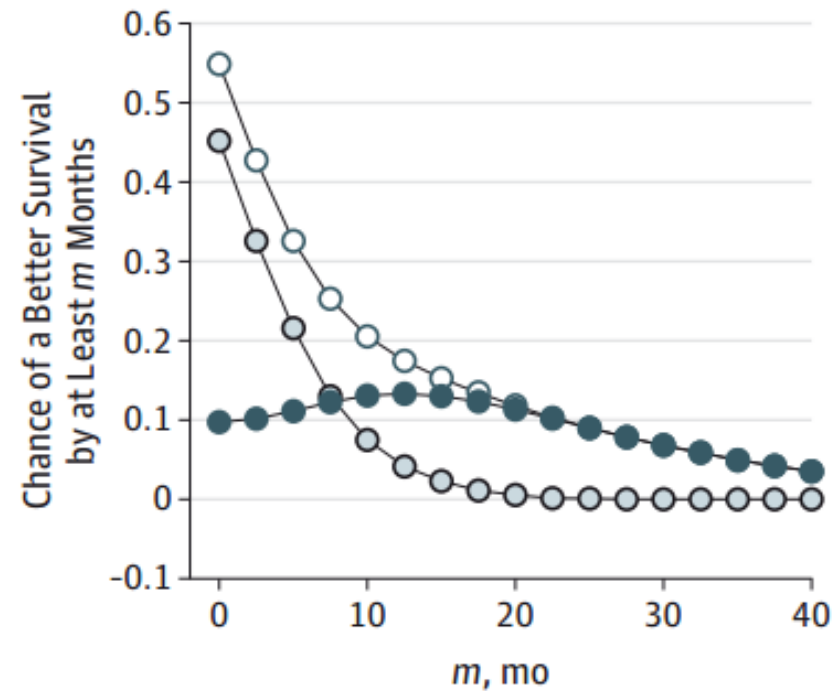
KAPLAN-MEIER CURVES

NET BENEFIT OF AT LEAST  $m$  MONTHS

**C** Scenario 3: delayed survival difference



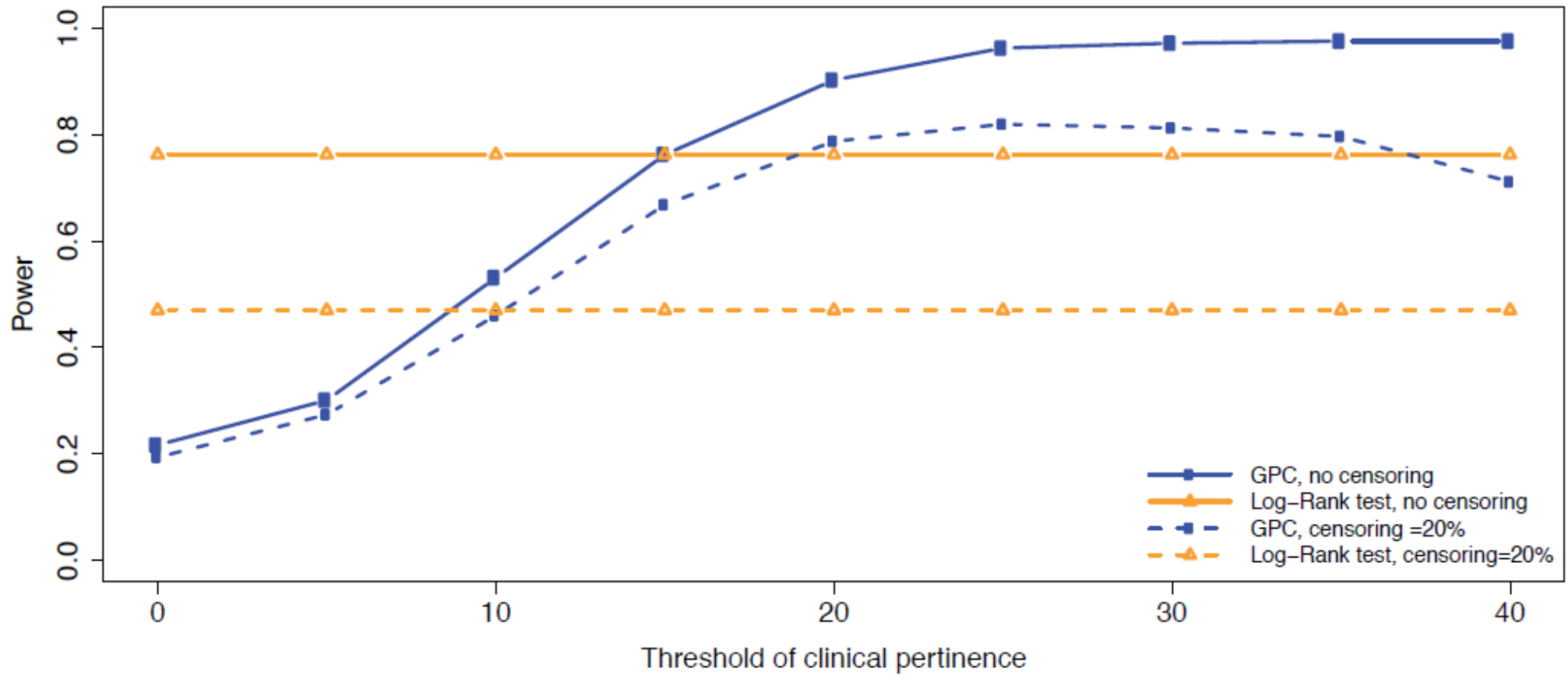
No. at risk	0	10	20	30	40
Group C	600	262	27	0	0
Group T	600	292	115	69	39



Example: immunotherapy for advanced solid tumors

# Power – Delayed Difference

Power of several tests in the delayed treatment effect scenario

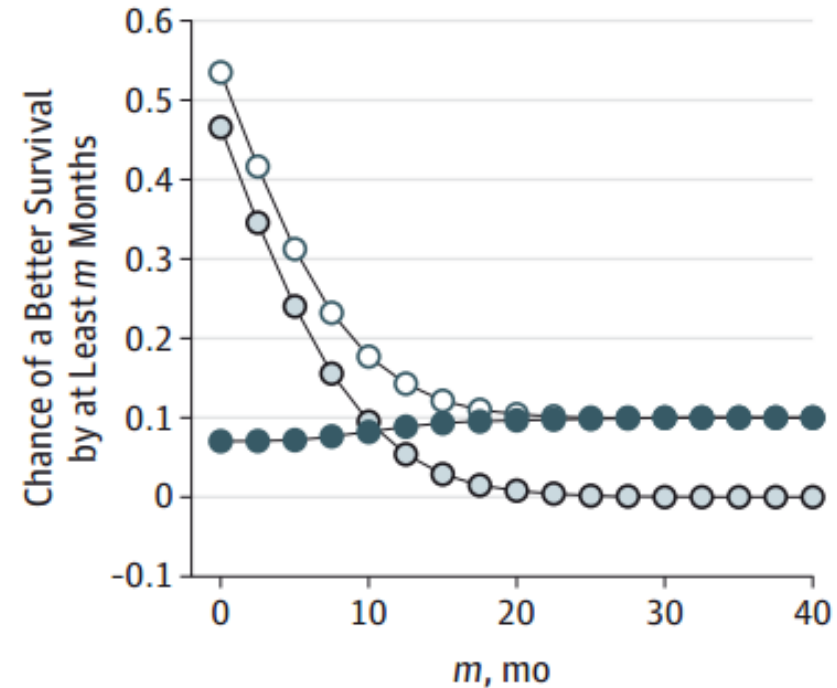
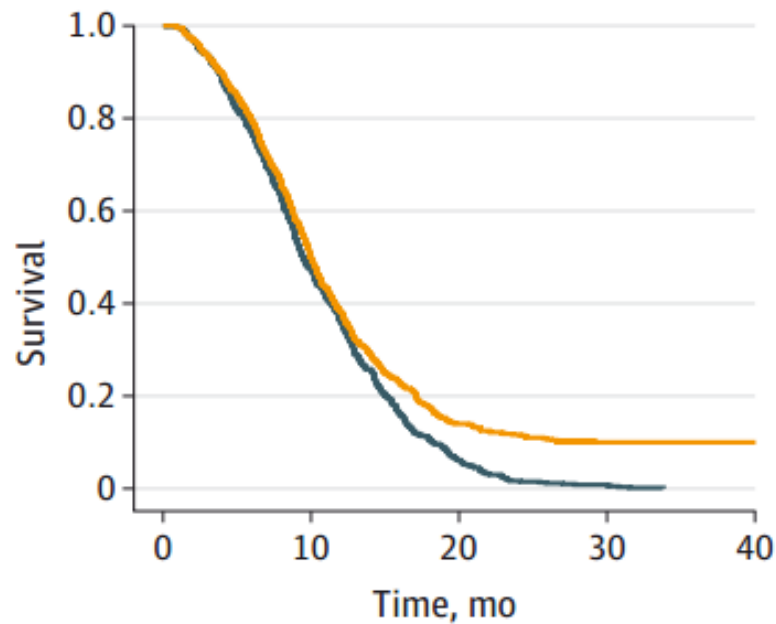


# Net Benefit – Cure Rate

KAPLAN-MEIER CURVES

NET BENEFIT OF AT LEAST  $m$  MONTHS

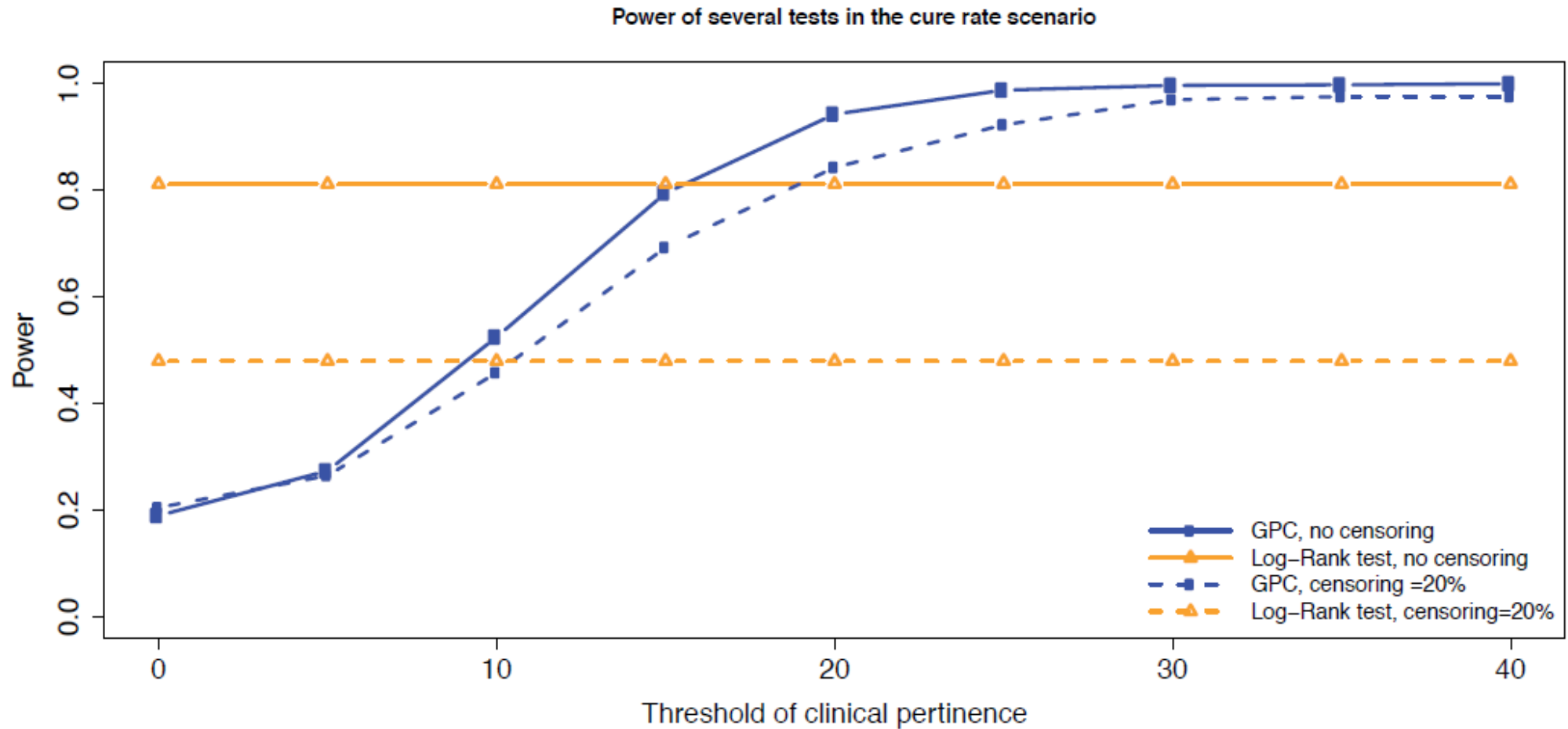
**D** Scenario 4: curable disease



No. at risk				
Group C	600	285	38	5
Group T	600	301	85	61

Example: allografts in childhood tumors

# Power – Cure Rate





- Related to the ‘probabilistic index’,  $P(X > Y)$ , [Acion et al. 2006, De Neve et al. 2013].
- The probabilistic index and related measures do not automatically generalize to other settings or different patient populations, where the variability of the outcome(s) of interest could be quite different [Senn 2011, Thas et al. 2012].
- These measures of benefit may be best seen as complementary to traditional (parametric) measures of benefit.

# Prioritized Outcomes

Now let  $(X_i, X'_i)$  and  $(Y_j, Y'_j)$  be observed values of two outcome measures, with  $X$  ( $Y$ ) being prioritized over  $X'$  ( $Y'$ )



$X_i$ vs. $Y_j$	$X'_i$ vs. $Y'_j$	Pair is
WIN		WIN
LOSS		LOSS
TIE or ?	WIN	WIN
TIE or ?	LOSS	LOSS
TIE or ?	TIE or ?	TIE or ?

- Assume  $K$  outcomes
- Let  $p_{ijk}$  ( $k=1, \dots, K$ ) be equal to
  - 1 if the pair is a win for the  $k$ -th outcome
  - 1 if the pair is a loss for the  $k$ -th outcome
  - 0 if the pair is a tie for the  $k$ -th outcome
- Define  $u_{ijk} = I(\text{the pair is ? for the } k\text{-th outcome})$
- Let
$$U(K) = \sum_{ij} \{p_{ij1}(1-u_{ij1}) + \dots + p_{ijK}u_{ij1} \dots u_{ij,K-1}(1-u_{ijK})\} / (nm)$$
- And

$$\Delta = E\{U(K)\}$$

## Myeloablative Autologous Stem-Cell Transplantation for Severe Scleroderma

N Engl J Med 2018;378:35-47.  
DOI: 10.1056/NEJMoa1703327

K.M. Sullivan, E.A. Goldmuntz, L. Keyes-Elstein, P.A. McSweeney, A. Pinckney, B. Welch, M.D. Mayes, R.A. Nash, L.J. Crofford, B. Eggleston, S. Castina, L.M. Griffith, J.S. Goldstein, D. Wallace, O. Craciunescu, D. Khanna, R.J. Folz, J. Goldin, E.W. St. Clair, J.R. Seibold, K. Phillips, S. Mineishi, R.W. Simms, K. Ballen, M.H. Wener, G.E. Georges, S. Heimfeld, C. Hosing, S. Forman, S. Kafaja, R.M. Silver, L. Griffing, J. Storek, S. LeClercq, R. Brasington, M.E. Csuka, C. Bredeson, C. Keever-Taylor, R.T. Domsic, M.B. Kahaleh, T. Medsger, and D.E. Furst, for the SCOT Study Investigators\*

### **METHODS**

We randomly assigned adults (18 to 69 years of age) with severe scleroderma to undergo myeloablative autologous stem-cell transplantation (36 participants) or to receive cyclophosphamide (39 participants). The primary end point was a global rank composite score comparing participants with each other on the basis of a hierarchy of disease features assessed at 54 months: death, event-free survival (survival without respiratory, renal, or cardiac failure), forced vital capacity, the score on the Disability Index of the Health Assessment Questionnaire, and the modified Rodnan skin score.

### **RESULTS**

In the intention-to-treat population, global rank composite scores at 54 months showed the superiority of transplantation (67% of 1404 pairwise comparisons favored transplantation and 33% favored cyclophosphamide,  $P=0.01$ ). In the per-protocol population

- Biostatistical Estimation of Net Effects For Individualization of Therapy
- Funds: the Walloon Region, Biowin – the Health Cluster of Wallonia and Innoviris, the Brussels Institute for Research and Innovation.
- Partners
  - International Drug Development Institute (IDDI)
  - Bristol-Myers Squibb
  - European Organization for Research and Treatment of Cancer (EORTC)
  - Université Catholique de Louvain (UCL)
  - Université Claude Bernard Lyon 1 (Lyon, France)



- Methods
  - Extensions of GPC (missing data, longitudinal, cross-over, ...)
  - “Optimal” GPC (censoring, risk/benefit, ...)
  - Comparisons with traditional methods
  - Use for trial design
- Applications (oncology, ophthalmology, ...)
- Software
  - Open
  - Proprietary (design, analysis, patient)

- Normally-distributed outcome  $Y$
- Two measurements: “earlier”  $Y_1$  and “later”  $Y_2$
- $Y_2$  “primary”
- For uncorrelated  $Y_1$  and  $Y_2$

$$\Delta = \theta_2 + \theta_1 * \left\{ \Phi\left(\frac{\mu_{20} + \tau_2 - \mu_{21}}{\sigma_2\sqrt{2}}\right) - \Phi\left(\frac{\mu_{20} - \tau_2 - \mu_{21}}{\sigma_2\sqrt{2}}\right) \right\}$$

$$= \{P(\text{T better for } Y_2) - P(\text{C better for } Y_2)\} + \\ \{P(\text{T better for } Y_1) - P(\text{C better for } Y_1)\} \cdot P(\text{tie on } Y_1)$$

$$= (\text{Net benefit for } Y_2) + (\text{Net benefit for } Y_1) \cdot P(\text{tie on } Y_2)$$

- Data for  $Y_1$  complete
- Data for  $Y_2$  missing completely at random in each treatment group
  - $\omega_0 = P(Y_2 \text{ observed for control})$ ,  $\omega_1 = P(Y_2 \text{ observed for treatment})$

- Then, for uncorrelated  $Y_1$  and  $Y_2$ ,

$$\Delta_{\text{MCAR}} = \omega_0 \omega_1 \Delta + \theta_1 (1 - \omega_0 \omega_1)$$

- Hence, estimation ignoring missing data (even for MCAR) is biased!

- A corrected estimator obtained from

$$\Delta = \{\Delta_{\text{MCAR}} - \theta_1 (1 - \omega_0 \omega_1)\} / (\omega_0 \omega_1)$$



- Formulae quickly complicate for correlated  $Y_1$  and  $Y_2$ 
  - MCAR:

$$\begin{aligned} \Delta = & \theta_2 \Phi(\beta_0) \Phi(\beta_0 + \gamma) + \theta_1 (1 - \Phi(\beta_0) \Phi(\beta_0 + \gamma)) + \\ & + \Phi(\beta_0) \Phi(\beta_0 + \gamma) \left( \frac{1}{\sigma_1^2} \int_{-\infty}^{\infty} \phi\left(\frac{y_{11} - \mu_{11}}{\sigma_1}\right) \left( BvN\left(h_{11}, \frac{y_{11} - \tau_1 - \mu_{10}}{\sigma_1}; \rho_3\right) - \right. \right. \\ & \left. \left. BvN\left(h_{12}, \frac{y_{11} - \tau_1 - \mu_{10}}{\sigma_1}; \rho_3\right) \right) dy_{11} - \right. \\ & \left. + \frac{1}{\sigma_1} \int_{-\infty}^{\infty} \phi\left(\frac{y_{10} - \mu_{10}}{\sigma_1}\right) \left( BvN\left(h_{13}, \frac{y_{10} - \tau_1 - \mu_{11}}{\sigma_1}; -\rho_3\right) - \right. \right. \\ & \left. \left. BvN\left(h_{14}, \frac{y_{10} - \tau_1 - \mu_{11}}{\sigma_1}; \rho_3\right) \right) dy_{10} \right) \end{aligned}$$

- Even more for MAR...
- Nevertheless, IPW estimators can be constructed

# Closing Remarks (1)

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- GPCs are attractive
  - In terms of patient centricity:
    - “Net benefit”, a patient-relevant measure
    - Accommodate prioritized outcomes
  - In statistical terms:
    - Equivalent to standard non-parametric tests in simple cases
    - May have better power than, e.g., the logrank test
    - Allow for testing of clinically relevant differences

# Closing Remarks (2)

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- GPCs require more fundamental research
  - In terms of theoretical properties:
    - sufficiency, completeness?
    - robustness to missing data
    - handling multiple relevance-thresholds
    - generalizability beyond the available sample?
    - ...
  - In terms of applicability:
    - Disease domains where additional insight can be obtained?